

1 Trend analysis by a piecewise linear regression model
2 applied to surface air temperatures in Southeastern Spain
3 (1973-2014)
4

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9

1 **Abstract:** The magnitude of the trends of environmental and climatic changes is mostly derived from the
2 slopes of the linear trends using ordinary least-square fitting. An alternative flexible fitting model,
3 piecewise regression, has been applied here to surface air temperature records in southeastern Spain for
4 the recent warming period (1973-2014) to gain accuracy in the description of the inner structure of
5 change, dividing the time series into linear segments with different slopes. Breakpoint years, with
6 confidence intervals (CIs), were estimated and separated periods of significant trend change were
7 determined. First, simple linear trends for mean, maximum and minimum surface air temperatures and
8 diurnal temperature range (DTR) from the four longest and most reliable historic records in SE Spain
9 were estimated. All series in the region showed intense linear warming signs during the period 1973-
10 2014. However, updated warming trends were lower than those previously cited for the region and Spain
11 from the 1970s onwards. Piecewise regression model allowed us to detect breakpoints in the series, and
12 the absence of significant trends in the most recent period of the segmented fits for two stations. In
13 general, piecewise regression model showed better fit than simple linear regression model, and thus,
14 showed a better description of temperature variability.

15

16 **Keywords:** temperature trends; piecewise regression; segmented regression; Southeastern Spain; regional
17 climate change; warming slow-down;

18

19

1 1. Introduction

2
3 Trend analysis is useful for a better understanding of climate change and variability. The estimation of
4 simple linear trends is the most straightforward assessment of the long-term behavior of a time series in
5 climate change studies. However, due to [short term](#) changes in the trends, real-world time series do not
6 generally fit in a straight line. Thus, simple linear regression may be misleading, as it does not describe
7 the inner structure of change in the series, ignoring the existence of significant changes in the slope of the
8 linear fit, called breakpoints. Such simple linear trends may be illogical and physically meaningless for
9 climatic data analyses, where the linearly fitted trend makes little sense, for the underlying mechanisms of
10 global climate change are likely to be nonlinear and nonstationary, so other methods of time series
11 analysis might be advisable (Seidel and Lanzante, 2004; Wu et al., 2007; Sefidmadgi et al., 2014). In
12 particular, linear trend does not adequately describe low-frequency behavior of temperature time series
13 (Karl et al., 2000). Piecewise regression model fits a nonlinear function with a non-constant rate of
14 change, and has been applied to analyze time series of different climatic variables to detect breakpoints in
15 linear trends. Karl *et al.*, (2000), identified the timing of change points in global temperature time series
16 by minimizing the residual sum of squares of all possible combinations of four line segments representing
17 time intervals of 15 years or more. Tome and Miranda (2004) adapted that fitting method to develop an
18 algorithm for fitting a continuous regression model with several break points to data and then it was
19 applied local changes in temperature, precipitations and the NAO index in Portugal. Liu et al. (2010) used
20 the same method to find partial trends of wind variability in the mesosphere and the lower thermosphere
21 over a local observatory at Collm, Germany.

22
23 Piecewise regression is a method of regression analysis where the response variable is split in two or
24 more intervals, and a line segment is fitted to each interval, with the constraint that the regression function
25 will be continuous. Each line is connected at an unknown value called breakpoint. Piecewise regression is
26 suitable for situations where the response variable shows abrupt changes within a few values of the
27 explanatory variable, (Toms and Lesperance, 2003). This flexible regression method is scarcely used in
28 the analysis of long term trends of climatic variables, though in many cases it offers a better fit to the
29 records, and shows better compliment with the assumptions of regression analysis.

30
31 Here we have applied to regional temperature records from SE Spain an alternative fitting approach for
32 long term climatic series, piecewise or segmented regression. SE Spain includes some of the most semi-
33 arid areas in the Northern Mediterranean, with increased vulnerability to projected shifts in global
34 circulation and pressure patterns, so it is of great interest the characterization and update of long-term
35 trends in this area. Climate change projections for the Mediterranean region show that high-temperature
36 conditions are generally expected to increase in the future (Giorgi and Lionello, 2008; Jacobeit et al.,
37 2014). Over the Iberian Peninsula, significant high rates of surface air temperature warming have been
38 recorded by the Spanish Meteorological Agency (AEMET) from the early 1970s onward, in accordance
39 with the last period of accelerated global warming (Hartmann *et al.*, 2013). For instance, Brunet et al.

1 (2007) estimated recent average rates of warming for Spain during 1973-2005 of $0.48\text{ }^{\circ}\text{C decade}^{-1}$ for
2 mean temperatures (T_{mean}), with 0.51 and $0.47\text{ }^{\circ}\text{C decade}^{-1}$ for maximum (T_{max}), minimum (T_{min})
3 temperatures respectively. Del Rio *et al.* (2012) reported rates of around $0.3\text{ }^{\circ}\text{C decade}^{-1}$ in over 90% of
4 Spanish weather stations for the period 1961-2006. However, most studies just have a limited spatial or
5 temporal coverage, and a scarce number of studies have estimated recent changes in long-term mean
6 annual temperature trends in Spain from the early 1970s into the 2010s (El Kenawy *et al.*, 2012; Turco *et*
7 *al.*, 2014; Gonzalez-Hidalgo *et al.*, 2015). In particular, assessments of observed changes in SE Spain are
8 scarce. Fernández-Montes and Rodrigo (2015) have recently reported temperature trends in SE Spain for
9 the period 1970-2007, estimating increases in T_{min} and T_{max} in most of the region of 1.2 and $0.5\text{ }^{\circ}\text{C}$
10 decade^{-1} respectively. However, no update of SE Spain temperature trends including recent 2010s years
11 has been reported to date.

12
13 Long-term global warming has been unequivocal from mid 1970s to 2013, with an average global trend
14 of $0.2\text{ }^{\circ}\text{C decade}^{-1}$ (Hansen *et al.*, 2010; Lawrimore *et al.*, 2011; Jones *et al.*, 2012; Rohde *et al.*, 2013;
15 Turco *et al.*, 2015). ~~However, recent warming trends throughout the planet have either decreased or lost~~
16 ~~statistical significance in many regional series in the planet from around 1997 (Easterling and Wehner,~~
17 ~~2009; Kaufmann *et al.*, 2011), though temperatures still remain well above the long-term average. Recent~~
18 ~~observations and global averages show a significant decrease in the warming trend from $0.12\text{ }^{\circ}\text{C decade}^{-1}$~~
19 ~~in the period 1951-2012 to $0.05\text{ }^{\circ}\text{C decade}^{-1}$ in 1998-2012 (Hartmann *et al.*, 2013). Nevertheless, such~~
20 ~~slowing for a decade or so has been seen in past observations and has been properly simulated in climate~~
21 ~~models (Met Office, 2013).~~ However, surface air temperatures are characterized by wide spatial and
22 temporal variability (Lovejoy, 2014; Steinmann *et al.*, 2015, Turco *et al.*, 2015), and there is a need to
23 determine local and sub-regional trends and variability to gain knowledge of global climate change
24 patterns.

25
26 A simple eye inspection of the main temperature records in SE Spain suggest a possible breakpoint year
27 at every series from the early 1990s (Fig. 1), showing decreased rates of warming since then. In order to
28 detect these breakpoints we have used a piecewise regression approach, estimating the continuous set of
29 straight lines that better fits every time series of annual T_{mean} , T_{max} , T_{min} , and diurnal temperature range
30 (DTR), deriving confidence intervals (CIs) for the breakpoint estimates. The main goals of this work
31 were: first use a simple linear regression model to update the long-term warming trends of air surface
32 temperatures at AEMET first order stations in SE Spain from the year 1973 to 2014; and second, to test
33 the goodness of fit and prediction skills of a piecewise regression model compared to simple linear
34 regression applied to the same period. ~~Our main interest in this work is to discuss if there is a recent~~
35 ~~significant breakpoint in long-term warming trends in SE Spain, and whether it is statistically consistent~~
36 ~~with the long-term warming, likely to happen superimposed on the longer-term warming trend, or whether~~
37 ~~the last records might represent a key emerging change point in long-term trends at the area. Here we~~
38 ~~present a statistical characterization of the recent changes in surface temperatures in this area, estimating~~
39 ~~both simple linear trends and main short-term inner changes of terms inside the period of study. Though~~

1 we suggest some possible physical explanations for some breakpoints, here we do not intend to present an
2 attribution analysis of the variability of the time series, though. Our study is limited to the detection and
3 descriptions of the patterns of change by regression analyses. In this sense, our objectives were:

4
5 1. Update temperature trends in SE Spain, based on the records of the main reliable stations and using
6 conventional linear regression fit.

7 2. Further describe the inner structure of temperature changes by fitting a nonlinear model (piecewise
8 regression)

11 **2. Data temperature series**

12
13 We have selected four meteorological stations with the longest, continuous and most reliable records in
14 SE Spain (Table 1): Almeria (AL), Granada (GR), Malaga (MA), and Murcia-San Javier (MU). These
15 stations are well-spaced across SE Spain (Fig. S1), with a minimum linear distance of 90 Km (between
16 GR and MA), and maximum of 340 km between MA and MU, below the threshold of 400 km that has
17 been suggested as optimal for building a representative meteorological network in Spain (Peña-Angulo et
18 al., 2014). AL, MA and MU are representative of coastal Mediterranean climate, while GR shows a
19 “continentalized” inland Mediterranean climate, with higher extremes of warm and cold days in summer
20 and winter, respectively. Their records cover at least from 1973 to 2014, including the recent period of
21 accelerated global warming from the 1970s that we intended to analyze in our area of study. These
22 stations are located within international airports, and belong to the first order (synoptic) network of the
23 Spanish official meteorological agency (AEMET). This selection was based on potential data quality
24 from highly monitored sites and good-quality records controlled by the *Servicio de Desarrollos*
25 *Climatológicos* (SDC, Climatological Branch) of AEMET. Raw data of daily T_{\max} and T_{\min} have
26 undergone quality control checks to avoid syntax, internal consistency and temporal coherence errors, and
27 controls of extreme thresholds and spatial coherency. Additionally, these records have been extensively
28 analyzed for artificial in-homogeneities in previous studies (Brunet et al., 2008; Staudt et al., 2007).

30 **3. Regression methods**

32 **3.1. Simple linear regression**

33
34 As a first conventional approach to detect temperature change, simple linear regression was applied to the
35 series. Trends and their 95% CIs were estimated by least squares linear regression. Linear trends were
36 estimated in every series from the slopes of the fit using values of annual averages of T_{\max} , T_{\min} , T_{mean} and
37 DTR, calculated from monthly means provided by AEMET.

39 **3.2. Piecewise regression**

1 |
2 However, when the residuals of simple linear fits for each T_{mean} series were tested for the assumptions of
3 normality, independence, homoscedasticity and linearity, it turned out that: a) homoscedasticity was not
4 met by GR station; b) the linear assumption was not verified by AL and MU residuals; c) the
5 independence assumption was rejected by Ljun-Box in AL. The violation of the homogeneous variance
6 assumption could result in unreliable estimates of the standard errors that might turn out in mistaken
7 conclusions over the slope. In these cases, heteroscedasticity-consistent (MacKinnon and White, 1985),
8 and autocorrelation-consistent estimators have been used (Newey and West, 1987). To solve these
9 problems, here we have tested an alternative regression model, piecewise regression. We have used a
10 segmented model between the mean response $E(Y)$ and the explanatory variable Z , modeled by adding in
11 the linear predictor the terms. Eq. (1):

$$\beta_1 z_i + \beta_2 (z_i - \psi)_+ \quad (1)$$

12
13
14 where β_1 is the slope of the left line segmented, β_2 is the ‘difference-in-slopes’, ψ is the breakpoint, and
15 $(z_i - \psi)_+ = (z_i - \psi) \times I(z_i > \psi)$ being $I(A)=1$ if A is true. In order to estimate the break-points
16 location, we use the approach suggested by Muggeo (2003) and at the R package ‘segmented’ (Muggeo,
17 2008). Karl *et al.* (2000) used this approach to obtain a better fit of global temperatures than simple linear
18 regression. Tomé and Miranda (2004) further developed an algorithm to identify best location for
19 breakpoints in climatic series. Here we have applied a piecewise regression model to those series where
20 there is enough statistical evidence to support the existence of breakpoints. Smoothed scatter plots were
21 used to provide the starting values for breakpoints in order to improve the convergence of the algorithm,
22 and we have checked the existence of a significant breakpoint by testing over the difference in slope
23 (Muggeo, 2003). We have employed the R package to estimate the parameters of the piecewise
24 regression in a deterministic way, and to fit linear segments to the data. The analysis of the residuals from
25 piecewise regression fits of our data showed that the assumptions of normality (Shapiro-Wilks and
26 Anderson-Darling tests), independence (Ljung-Box test) and homoscedasticity (Breusch-Pagan test) were
27 met by all the series, with the exception of GR. In this case, the robust variance estimator proposed by
28 MacKinnon and White (1985) was used. In order to test for a significant slope, we have applied a Wald’s
29 test. As general criteria, we have not estimated the slope of segments when they represented time
30 intervals of less than 5 years. As stated by Tome and Miranda (2005), “, if the first or the last breakpoint
31 happen near the minimum allowed position the result should be looked upon with some suspicion”. These
32 graphs are not shown in the figures at the results section. hh

33 34 35 **3.3. Evaluation of regression models**

36
37 We have compared both fitting and predicting performance of piecewise and simple linear models applied
38 to the whole length of data available for every series of T_{mean} , T_{max} , T_{min} and DTR (Table 2). In order to

1 compare the goodness-of-fit of the models, we calculated R^2 for each fit, and the residual standard error
2 (RSE) in order to avoid the artificial skill of R^2 . Also, we carried out a cross-validation analysis (Hastie et
3 al., 2009) to compare the goodness-of-forecasting skills among models. Thus, the data were divided into 5
4 roughly equal-sized parts, and for the i_{th} part, $i=1, \dots, 5$, the model was fitted using the other 4 parts of the
5 data. The prediction error of the fitted model was estimated when predicting the i_{th} part of the data.
6 Finally, the mean square error (MSE_{CV}) of the 5 evaluation parts was calculated as forecasting skill
7 indicator. As we can observe in Table 2, piecewise regression showed a superior behavior both fitting and
8 forecasting compared to simple linear regression. It must also be noted that, due to smoothing by
9 averaging extreme temperature values, the differences between piecewise and linear regression methods
10 are less in T_{mean} than in T_{max} or T_{min} .

11
12 In general, piecewise regression allowed obtaining a better fit and an improved predictive performance
13 than simple linear model. Though it can be argued that R^2 values are affected by a higher number of
14 parameters used for piecewise model, the values of adjusted R^2 , where the effect of the number of
15 parameters is removed, also confirmed the better fit for piecewise regression. Furthermore, in cases where
16 linear regression provided a very poor fitting, as with T_{max} in AL adjusted $R^2=0.001$, the piecewise
17 regression yields a fairly good fitting (adjusted $R^2=0.49$). Better performance is also characterized by
18 lower residual standard error (RSE), which is not affected by the differences in number of parameters.
19 These RSE are lower in piecewise regression than linear regression (Table 2) (except in T_{min} for MA).
20 Hence, we believe the differences in the breakpoints found in nearby stations are not a side effect of the
21 statistical model used, but they could be instead explained by either undetermined local forcings or
22 internal variability.

23 24 **4. Results and discussion**

25
26 A casual eye inspection of SE Spain temperature series for 1973-2014 suggests two periods of different
27 change rates (Fig. 1): a first period of intense warming, from 1973 until the early 1990s; and a second
28 period of lower or not significant rates of temperature increase from 1989 to 2014. ~~The sharp drop in~~
29 ~~1991, associated to Mt. Pinatubo eruption, was transient and widespread throughout the planet.~~ As a first
30 approach, simple linear regression slopes for 1973-2014 showed average warming trends of T_{mean} in the
31 range of 0.25-0.46 $^{\circ}C$ decade $^{-1}$ (Table 3), below average rates cited in previous studies for this area and
32 for Spain (Brunet *et al.*, 2007; Fernandez-Montes and Rodrigo, 2015). However, these studies don't cover
33 the most recent records of the 2010s, and they are difficult to compare with our results because of the
34 different areas, data sets and periods used (Brunet *et al.*, 2007; Del Rio *et al.*, 2012; Gonzalez-Hidalgo *et*
35 *al.*, 2015). In all stations except GR, mean annual DTR decreased at significant rates, between -0.25 and-
36 0.36 $^{\circ}C$ decade $^{-1}$, due to higher increases of T_{min} (0.43/0.58 $^{\circ}C$ decade $^{-1}$) than T_{max} . This trend in DTR is in
37 agreement with the expected "fingerprint" for global greenhouse warming (Hartmann *et al.*, 2013), but is
38 opposite to the positive trend in annual DTR reported for Spain for the last decades (Galan *et al.*, 2001;
39 Brunet *et al.*, 2007; Del Rio *et al.*, 2012). However, DTR reductions have been cited in Southern parts of

1 Spain (Horcas *et al.*, 2001; Staudt *et al.*, 2005; Gonzalez-Hidalgo *et al.*, 2015) and have been related to
2 T_{\min} increases driven by land-use changes. The physical interpretation of DTR trends, and the
3 contribution of local forcings due to land-use changes still remain a key uncertainty, and confidence in
4 this indicator has been defined as *medium* in reported decreases in observed DTR (Hartmann *et al.*, 2013).
5 Spatial variability of trends in T_{\min} and T_{\max} was high, as described by with Peña-Angulo *et al.* (2014),
6 who detected in SE coastal Mediterranean areas the highest variability for mainland Spain, and again
7 suggested an association with the dramatic land-use changes by urbanization and agriculture at this coast
8 in the last decades. However, the two more distant stations, MA and MU, two coastal stations that lie 340
9 km away, showed comparable patterns of change for all variables, different to trends in GR and AL, with
10 higher rates of increase in all variables, and T_{\min} rates of increase of $0.5\text{ }^{\circ}\text{C decade}^{-1}$. In GR no significant
11 change in long-term DTR was detected, with only slight differences in rates of warming for T_{\max} and T_{\min} ,
12 in agreement with described patterns from 1970s for Spanish southern plateau (Galan *et al.*, 2001), with a
13 continentalized Mediterranean climate that is more similar to GR climate than to coastal stations. AL
14 showed the lowest rates of warming in the area, probably related to the dramatic land-cover change
15 towards greenhouse farming developed from the 1980s, resulting in the widest concentration of high
16 albedo greenhouses in the planet (Fig. S2). The alteration of the local energy balance by widespread
17 reflectivity increase has probably offset in this area the global forcing exerted by the increase of
18 greenhouse gases (Campra *et al.*, 2008; Campra and Millstein, 2013). DTR decrease due to higher rates of
19 warming of T_{\min} than T_{\max} has been also associated to urbanization and land cover changes (Mohan and
20 Kandya, 2015). MA is the only station in this study that might have been affected by the rapid growth of
21 urban areas (Malaga city), and the recent growth of low albedo pavements around it by recent
22 development of taxiways and runways from 2005 (Fig. S3) (Fomento Ministry of Spain, 2007). This
23 recent land cover change around the station might have enhanced warming trends in this record,
24 increasing heat retention properties of the surface (McKittrick and Michaels, 2007). In GR and MU, little
25 increase in urbanization has been produced around the stations, located in international airports far from
26 the growing cities, so trends in temperatures and DTR cannot be clearly associated here to local land
27 cover changes. Besides these local factors, the influence of the Mediterranean is a key forcing of SE
28 Spain temperature records. Observations of Mediterranean surface temperature indicate strong warming
29 from 1973. The average rate of warming at Western Mediterranean was $0.22\text{ }^{\circ}\text{C decade}^{-1}$ over 1973-2008
30 (Skirlis *et al.*, 2012). The Alboran Sea surrounding SE Spain showed an averaged warming trend of 0.31
31 C decade^{-1} on 1982-2012 (Shaltout and Omstedt, 2014).

32

33 Alternatively, piecewise regression model from the full historic records available at each station offered a
34 better fit to the data in all temperature series (Fig. 2), showing that this model represents the observations
35 more accurately than the simple linear regression model generally used to describe climatic trends. As a
36 general criteria, in figures 2, 3 and 4 we have not represented piecewise regressions where last segment
37 are less than 5 years long. Uncertainty analysis of piecewise regression fits for every series and variable
38 are given in Tables S1-S4. This flexible regression model allowed for detection of several trend change
39 points, located in different years for every series and for every variable (Table 4). Recent T_{mean} trends

1 were not significant from these breakpoints in AL and GR. In MU, T_{mean} increase was less intense since
2 the 1982 breakpoint (Table 5). On the contrary, a breakpoint in 2013 was found in T_{mean} in MA driven by
3 the historic T_{mean} record in 2014 (Table 4).

4
5 In order to identify the components of change of T_{mean} , we have carried out additional piecewise
6 regression fits to T_{max} and T_{min} (Table 6). Breakpoints in both T_{max} and T_{min} were detected for all series
7 except MA (Fig. 3). DTR showed decreasing trends in the last segments of AL and MU (Fig. 4), but
8 recent breakpoints towards increasing DTR have appeared in MA (2012) and GR (2011), though only
9 significant at 10% (figures not shown) (Table 6). However, we found no signs in our data of the pattern of
10 change in described in global DTR, characterized by a decline and subsequent increase from a breakpoint
11 in mid 1980s, associated to global dimming and subsequent brightening (Wild *et al.*, 2007). (Table 6). In
12 AL, the 1989 breakpoint in T_{mean} was driven by a significant fall of T_{max} from 1987, further maintained by
13 the later stabilization of T_{min} (with no significant trend from 1998). As a consequence of these trends, AL
14 station showed the highest recent rates of DTR reduction in the area ($-0.61\text{ }^{\circ}\text{C decade}^{-1}$ since 1982),
15 probably due to local land cover changes stated before. In GR, the 1997 breakpoint in T_{mean} was driven by
16 stabilization of T_{min} , with no significant trend from 1997. In MU a breakpoint was located around the
17 early 1980s in both T_{max} and T_{min} , coincident with T_{mean} breakpoint. No significant breakpoints in T_{max} and
18 T_{min} were detected in MA, showing a continuous warming trend during the period of study, along with the
19 highest simple linear rates in the region ($+0.46\text{ }^{\circ}\text{C decade}^{-1}$). MA is the only series that shows no clear
20 signs of warming slow-down in T_{min} and T_{max} in recent years. Furthermore, a 2013 breakpoint was
21 detected in T_{mean} , ~~due to 2014 breaking record,~~ increasing long-term rates of warming. A similar 2013
22 breakpoint was shown by T_{max} in GR. ~~Our regression analyses does not allow us to suggest that these~~
23 ~~recent breakpoints are signs of a future increase in warming trends in these stations, and this might~~
24 ~~lead to failed conclusions when piecewise segments length is not long enough (Karl *et al.*, 2000).~~

25 26 5. Conclusions

27
28 Simple linear trends of temperatures at SE Spain for the period 1973-2014 reported here are consistent
29 with the magnitude of warming described previously for the region. However, our results show that the
30 generalized use of simple linear regression fits for the estimation of long term trends might not be
31 sufficiently accurate to describe the structure of change of temperatures in long-term series, while flexible
32 models such as piecewise regression provide better fits, and allows the detection of key breakpoints in the
33 trends. Besides, the method is also much simpler and easier to interpret than most non-linear function
34 fitting. We believe that our analysis of piecewise regression fit shown here justifies the suitability of the
35 employment non-linear models in climatic change studies. Results obtained with our piecewise regression
36 model showed that it can be a useful complement to the conventional simple linear models for a more
37 detailed description of changes in climatic variables. By flexible regression, we have detected a recent
38 slow-down in long-term warming trends in three of the four main records in SE Spain, with no significant
39 trends in T_{mean} in the most recent segments of two of them (AL and GR) ~~no matter the uncertainty in the~~

1 ~~exact location of breakpoints reported here, given by 95% CIs (Table 4), we have shown statistical~~
2 ~~evidence of the presence of such reduction in warming rates in the area, consistent with previous reports~~
3 ~~of recent decreasing warming trends in global series.~~ However, and given the limited length of the
4 observations, we cannot suggest by our analysis that the absence of statistically significant warming in the
5 most recent segments of piecewise regression model are necessarily inconsistent with historic variability
6 of the series studied. ~~Also it is important to notice that by simply using both regression analysis, linear~~
7 ~~and piecewise, we cannot make projections of future trends, neither suggest that reduced linear trends in~~
8 ~~the area of study continue in the next years.~~ Furthermore, recent breakpoints observed in MA and GR,
9 ~~might be signing increasing warming trends in the future, enhanced by 2014 global historic record~~
10 ~~(Jones, 2015), year-to-date global temperatures (up to June 2015) (NOAA, 2015a).~~ Spatial
11 differences in long-term trends and breakpoints location might be due to undetermined local dynamics in
12 climate and/or forcings that still remain to be determined in every station, or just simply due to natural
13 variability. ~~It must also be taken into account that the method of piecewise regression is strongly~~
14 ~~dependent on the first guesses and imposed parameters such as minimum time distance between~~
15 ~~breakpoints, and minimum trend change between breakpoints (Tome and Miranda, 2005). Future~~
16 ~~attribution studies, such as those recommended at IPCC Good Practices Guidance (Hegerl et al., 2010),~~
17 ~~including numerical simulations with mesoscale models and extended local forcings might help~~
18 ~~explaining the observed variability, and the location of the trend changing points in the temperature series~~
19 ~~described here. Complementary to this search of statistical relationships, recent simulation studies with~~
20 ~~optimized treatment of global climate forcings have shown better insight to possible causes of a possible~~
21 ~~slow-down in global warming (Smith et al., 2014).~~

22
23 **Author contribution.** P. Campra designed the scope of the research, performed the simple linear
24 regressions and wrote the manuscript, and M. Morales carried out all deep statistical analyses, residuals
25 tests and developed the piecewise regression model in R.

26
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31 32 **Supporting Information**

33
34 Map of location of stations in SE Spain, picture of greenhouses land cover around AL station, picture of
35 MA airport new taxiways. Tables of uncertainty estimates of piecewise regression of mean, maximum,
36 minimum and DTR temperature series. This material is available free of charge via the Internet at ...

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1 **Table 1. First order Spanish Meteorological Agency (AEMET) stations in SE Spain. *Length of 2-m**
 2 **temperature records available.**

	ALMERIA (AL)	GRANADA (GR)	MALAGA (MA)	MURCIA (MU)
Location	36° 50' 47" N 2° 21' 25" W	37° 11' 23" N 3° 47' 22" W	36° 39' 58" N 4° 28' 56" W	37° 47' 20" N 0° 48' 12" W
Altitude (msl.)	21	567	5	4
Length*	1972-2015	1973-2015	1950-2015	1946-2015

3

4

1 **Table 2. Goodness-of-fit test (R^2 /RSE), and predictive value (MSE_{CV}) of the two regression models for T_{mean} ,**
 2 **T_{max} , T_{min} y DTR. R^2 = coefficient of determination; Adj. R^2 = adjusted coefficient of determination; RSE=**
 3 **residual standard error; MSE_{CV} = mean square error for cross-validation.**

		PIECEWISE				SIMPLE LINEAR			
		R^2	<u>Adj. R^2</u>	RSE	MSE_{CV}	R^2	<u>Adj. R^2</u>	RSE	MSE_{CV}
T_{mean}	AL	0.60	<u>0.58</u>	0.32	0.14	0.40	<u>0.38</u>	0.39	0.14
	GR	0.42	<u>0.39</u>	0.53	0.29	0.36	<u>0.34</u>	0.55	0.32
	MA	0.76	<u>0.75</u>	0.33	0.12	0.74	<u>0.73</u>	0.34	0.12
	MU	0.67	<u>0.65</u>	0.38	0.15	0.59	<u>0.58</u>	0.42	0.19
T_{max}	AL	0.52	<u>0.50</u>	0.37	0.17	0.03	<u>0.001</u>	0.52	0.27
	GR	0.39	<u>0.35</u>	0.64	0.32	0.28	<u>0.27</u>	0.68	0.33
	MU	0.57	<u>0.55</u>	0.31	0.10	0.49	<u>0.48</u>	0.33	0.12
T_{min}	AL	0.70	<u>0.69</u>	0.36	0.17	0.67	<u>0.66</u>	0.38	0.27
	GR	0.40	<u>0.37</u>	0.58	0.38	0.29	<u>0.27</u>	0.62	0.45
	MA	0.79	<u>0.78</u>	0.38	0.30	0.79	<u>0.78</u>	0.37	0.26
	MU	0.66	<u>0.64</u>	0.53	0.28	0.58	<u>0.57</u>	0.58	0.38
DTR	AL	0.73	<u>0.72</u>	0.35	0.16	0.46	<u>0.44</u>	0.49	0.24
	GR	0.21	<u>0.17</u>	0.62	0.38	0	<u>-0.02</u>	0.69	0.47
	MA	0.53	<u>0.50</u>	0.31	0.11	0.49	<u>0.47</u>	0.32	0.11
	MU	0.46	<u>0.43</u>	0.42	0.18	0.41	<u>0.39</u>	0.43	0.20

4

5

1 **Table 3. Change in annual mean (T_{mean}), maximum (T_{max}), minimum temperatures (T_{min}), and diurnal**
 2 **temperature range (DTR), estimated as the slope of a simple linear regression fit (in $^{\circ}\text{C decade}^{-1}$), and**
 3 **associated 95% confidence intervals (CI) for the recent warming period 1973-2014, at SE Spain first order**
 4 **AEMET stations. Not significant values in italics ($p>0.05$).**

Station	T ^a series	Decadal coefficient ($^{\circ}\text{C}$)	95% CI
AL	T_{mean}	0.25	(0.15/0.36)
	T_{max}	<i>0.07</i>	<i>(-0.07/0.2)</i>
	T_{min}	0.43	(0.34/0.53)
	DTR	-0.36	(-0.49/-0.24)
GR*	T_{mean}	0.33	(0.19/0.48)
	T_{max}	0.34	(0.17/0.52)
	T_{min}	0.32	(0.16/0.48)
	DTR	-0.02	<i>(-0.16/0.20)</i>
MA	T_{mean}	0.40	(0.30/0.51)
	T_{max}	0.26	(0.18/0.35)
	T_{min}	0.55	(0.40/0.70)
	DTR	-0.29	(-0.4/-0.18)
MU	T_{mean}	0.46	(0.37/0.55)
	T_{max}	0.34	(0.24/0.43)
	T_{min}	0.58	(0.48/0.68)
	DTR	-0.25	(-0.34/-0.17)

5

6

1 **Table 4. Breakpoint estimates and 95% CIs for annual T_{mean} , T_{max} , T_{min} and DTR series in SE Spain. Values in**
 2 ***italics* are not significant at 5% level. ⁺ Significant at 10%**

Station	T_{mean}	T_{max}	T_{min}	DTR
AL	1989 \pm 5.3	1987 \pm 3.4	1998 \pm 11.4	1982 \pm 2.6
GR	1997 \pm 11.8	2013 \pm 1.0	1997 \pm 9.2	2011 \pm 2.3
MU	1982 \pm 5.7	1983 \pm 7.2	1981 \pm 5.0	1981 \pm 4.1 ⁺
MA	2013 \pm 1.2 ⁺	-	<i>1977 \pm 10.7</i>	2012 \pm 3.2 ⁺

3

4

1 **Table 5. Trends in °C decade⁻¹ of annual mean temperature change (T_{mean}) from simple linear regression fit**
 2 **(SL) for each station (in bold), compared to the two successive sub-periods defined by piecewise (PW)**
 3 **regression. Trend values in *italics* are not significant at 5% level. *Breakpoint was detected in 2013 and no**
 4 **segmented regression periods are reported.**

Station	Regression model	Period	Estimate	95% CI
AL	SL	1973-2014	0.25	(0.15/0.36)
	PW ₁	1973-1989	0.76	(0.51/1.0)
	PW ₂	1989-2014	<i>-0.07</i>	<i>(-0.15/0.14)</i>
GR	SL	1973-2014	0.33	(0.19/0.48)
	PW ₁	1973-1997	0.55	(0.34/0.76)
	PW ₂	1997-2014	<i>-0.03</i>	<i>(-0.57/0.51)</i>
MU	SL	1973-2014	0.40	(0.30/0.51)
	PW ₁	1973-1982	1.28	(0.69/1.87)
	PW ₂	1982-2014	0.27	(0.14/0.4)
MA*	SL	1973-2014	0.46	(0.37/0.55)

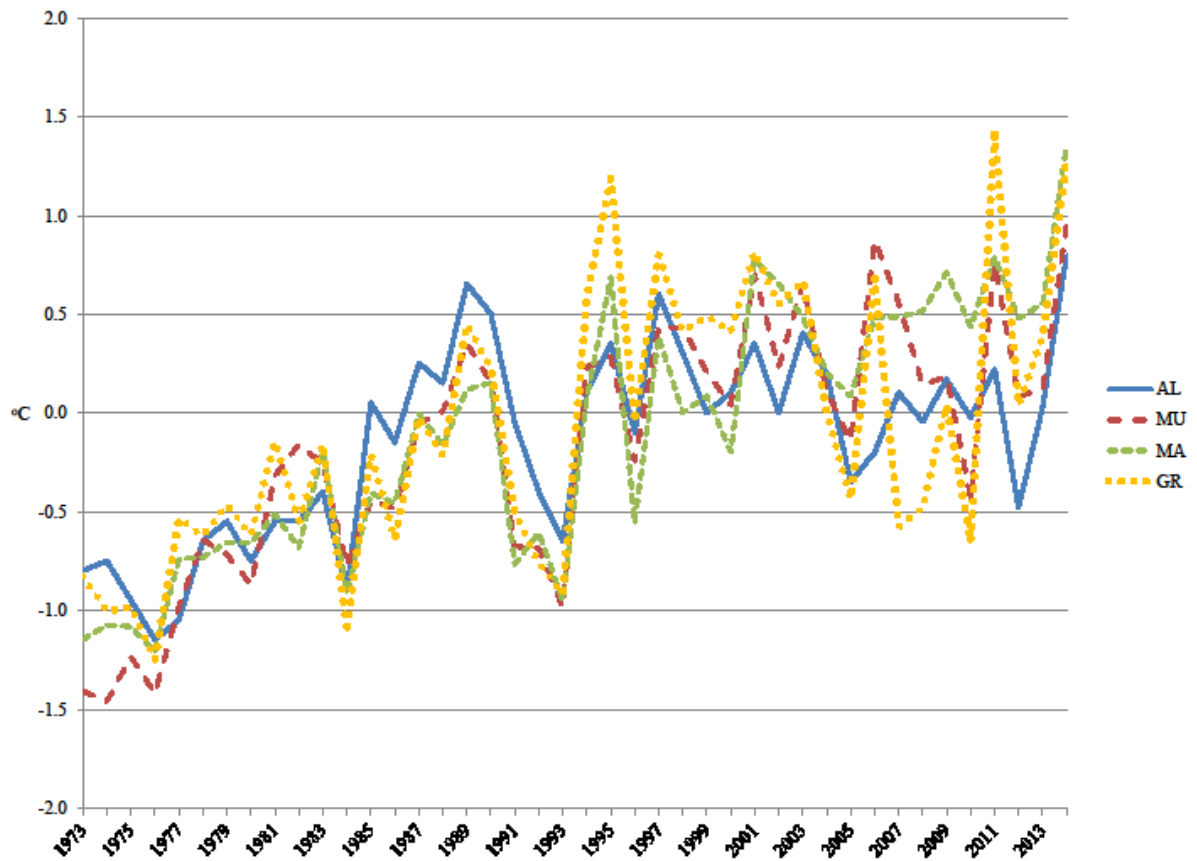
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1 **Table 6. Trends in °C per decade of annual maximum (T_{\max}) and minimum (T_{\min}) temperature anomalies and**
 2 **diurnal temperature range (DTR) for the periods defined by piecewise (PW_{1,3}) regression model. Trend values**
 3 **in *italics* are not significant at 5% level. + Significant at 10%**

Station		T_{\max}	T_{\min}	DTR
AL	Breakpoint	1987	1998	1982
	PW ₁	1.08	0.58	1.31
	PW ₂	-0.32	<i>0.16</i>	-0.61
GR	Breakpoint	2013	1997	2011
	PW ₁	0.29	0.63	<i>-0.09</i>
	PW ₂	-	<i>-0.18</i>	-
MU	Breakpoints	1983	1981	1981⁺
	PW ₁	0.82	1.92	-1.01
	PW ₂	0.16	0.38	<i>-0.20</i>
MA	Breakpoint	-	-	2012⁺
	PW ₁	-	-	-0.28
	PW ₂	-	-	-

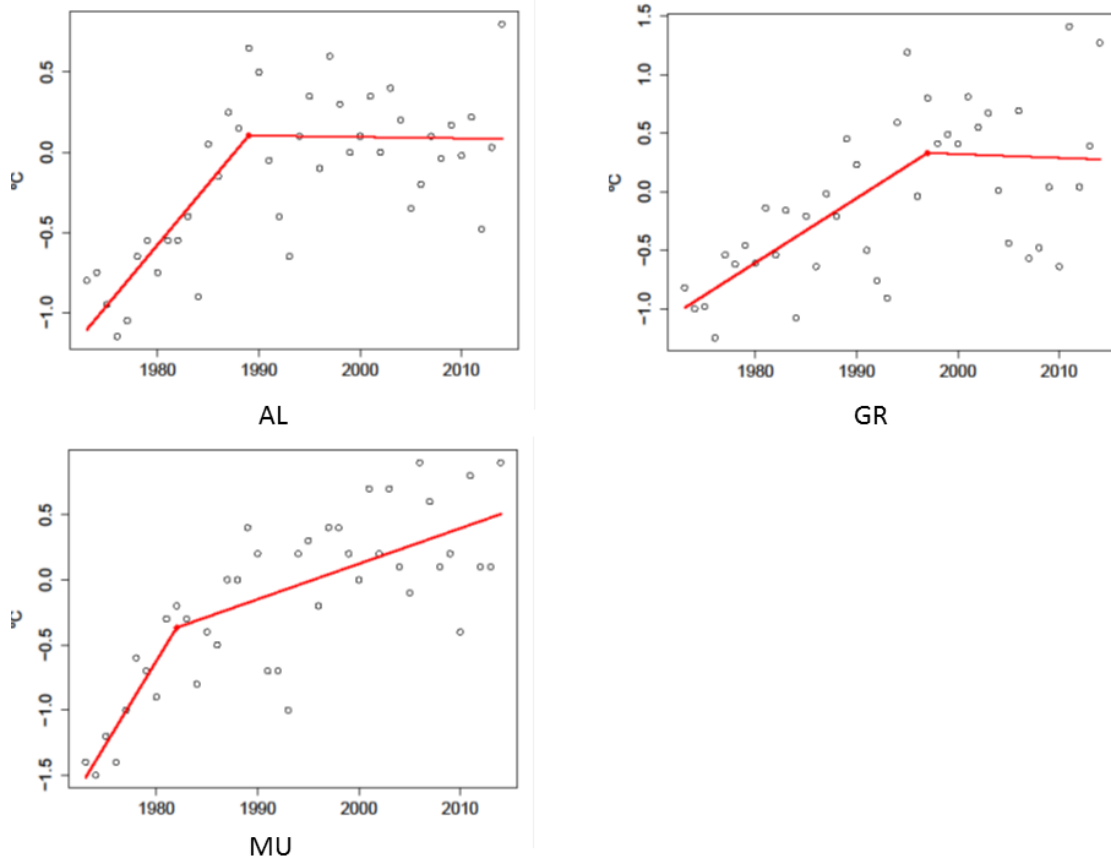
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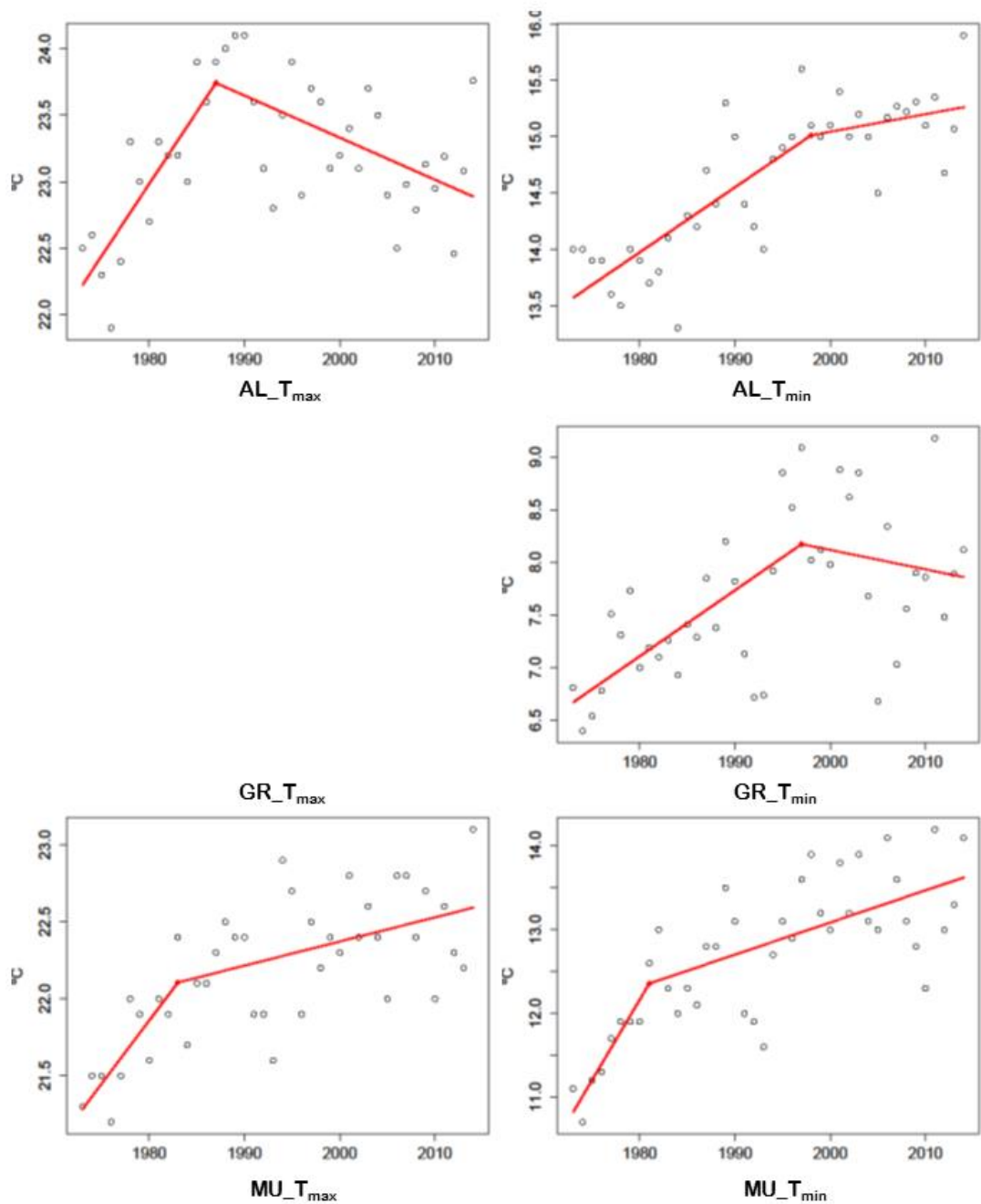
3 **Figure 1. Mean annual temperature anomaly series (°C) from 1973-2014 at SE Spain First order stations of the**
 4 **Spanish official meteorological network (AEMET). Reference period: 1981-2010. AL=Almeria; GR=Granada;**
 5 **MA=Malaga; MU=Murcia**

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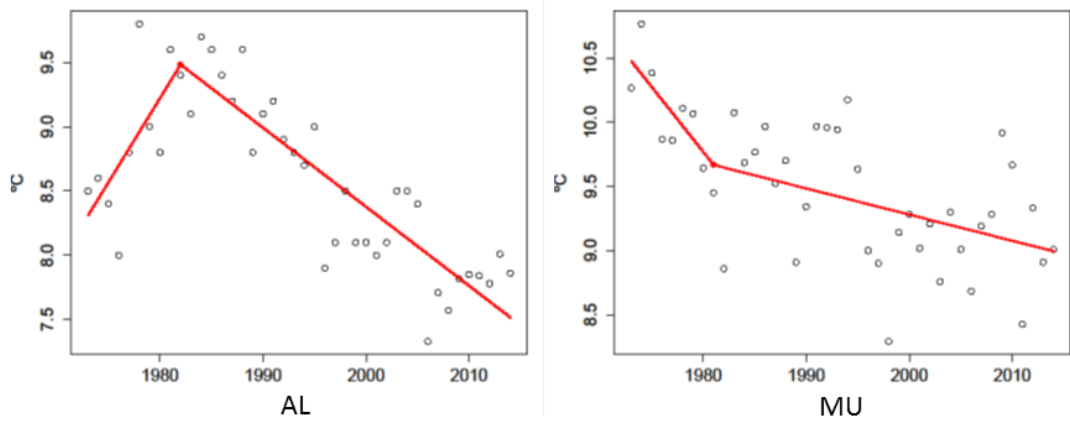
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Figure 2. Piecewise regression fitting of historic records of mean annual temperature anomalies (°C) in SE Spain. Only those series with last segment > 5 years are shown. AL=Almeria; GR=Granada; MU=Murcia.



1

2 **Figure 3. Piecewise regression fitting of historic records of annual average maximum (Tmax) and minimum**
 3 **(Tmin) temperature anomalies (°C) in SE Spain stations. Only those series with last segment > 5 years are**
 4 **shown. AL=Almeria; GR=Granada; MU=Murcia.**



1

2

3 **Figure 4. Piecewise regression fitting of historic records of diurnal temperature range (DTR) in °C in SE Spain**

4 **stations. Only those series with last segment > 5 years are shown. AL=Almeria; MU=Murcia. (MA and GR**

5 **series not shown)**

6