

Multifractal behaviour of the soil water content of a vineyard in NW Spain during two growing seasons

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Abstract

Soil processes are characterized by a great degree of heterogeneity, which may be assessed by scaling properties. The aims of the current study were to describe the dynamics of soil water content at three depths in a vineyard under rain-fed and irrigation conditions and to assess the multifractality of these time data series. Frequency domain reflectometry (FDR) sensors were used for automatically monitoring soil water content in a vineyard located in Leiro (Ourense, NW Spain). Data were registered at 30-minute intervals at three depths (20, 40 and 60 cm) between 14th June and 26th August 2011 and 2012. Two treatments were considered: rain-fed and irrigation to 50% crop evapotranspiration. Soil water content data series obeyed power laws and tended to behave as multifractals. Values for entropy (D_1) and correlation (D_2) dimensions were lower in the series from the irrigation treatment. The Hölder exponent of order zero (α_0) was similar between treatments; however, the widths of the singularity spectra, $f(\alpha)$, were greater under irrigation conditions. Multifractality indices slightly decreased with depth. These results suggest that singularity and Rényi spectra were useful for characterizing the time variability of soil water content, distinguishing patterns among series registered under rain-fed and irrigation treatments.

1

2 **1 Introduction**

3 Soil water storage variability is strongly related with topographical, geological, edaphic and
4 vegetation factors (Braud et al., 1995). These environmental factors and processes (rainfall,
5 evapotranspiration, runoff) do not operate independently but as a conjunction of processes
6 with nested and complex effects. Overall, this results in a distribution of soil water storage
7 that varies as a function of the temporal and spatial scales. Therefore, similar to other soil
8 properties and processes (Western and Blöschl, 1999; Zeleke and Si, 2006), soil water storage
9 along time is a complex process characterized by a lack of homogeneity; heterogeneity in
10 space and/or time is a feature that can be described by scaling procedures.

11 Fractals have been widely employed in soil science, as soil properties may be described
12 through scale invariance concepts (Tyler and Wheatcraft, 1990; Perfect et al., 1996; Vidal
13 Vázquez et al., 2007; Biswas et al., 2012a). More recently, several authors performed
14 multifractal studies of heterogeneous time data series. For instance, Jiménez-Hornero et al.
15 (2010) described ozone time series using the multifractal formalism. Rodríguez-Gómez et al.
16 (2013) used a multifractal approach for characterizing solar radiation time series.

17 Soil water content can be automatically estimated by using sensors that measure variations in
18 the soil dielectric constant, since it is strongly related with soil water content (Mestas-Valero
19 et al., 2012). This parameter is characterized by its spiky dynamics, with sudden and intense
20 peaks of high frequency activity, mostly at soil surface. Several studies have described scaling
21 patterns for the behaviour of soil water content spatial distribution (e.g. Kim and Barros, 2002;
22 Biswas et al., 2012b); however, multifractal analyses of continuously measured soil water
23 content are scarce, except for a study on rain-fed grassland (Mestas-Valero et al., 2011).
24 Therefore, the aim of the current work was to describe soil water dynamics in a vineyard
25 subjected to two different treatments (rain-fed and irrigated) and to assess multifractality of
26 these data series over two consecutive seasons.

27

1 **2 Materials and Methods**

2 **2.1 Description of the study area**

3 The experiment was conducted over two consecutive growing seasons (2011-2012) in a 0.2-
4 ha vineyard (*Vitis vinifera* L.) planted with cultivar ‘Albariño’, located in the experimental
5 farm of the Estación de Viticultura e Enoloxía de Galicia (EVEGA), in Leiro (42° 21.6’ N, 8°
6 7.02’ W, elevation 115 m), Ourense, Spain (Fig. 1). Vines were grafted in 1998 on 196-17C
7 rootstock and trained to a vertical trellis on a single cordon system (10-12 buds per vine).
8 Rows were east-west oriented, spacings between vines and between rows were 1.25 and 2.4
9 m, respectively (3333 vines ha⁻¹). The soil at the site was sandy-textured (64% sand, 16% silt,
10 20% clay), slightly acidic (pH 6.3), medium fertility (2.7% organic matter) and with a rather
11 shallow profile (≈1.2 m). The climate of the studied site is temperate, humid with cool nights
12 (Fraga et al., 2014).

13 **2.2 Experimental design**

14 The reference evapotranspiration (ET₀) per week for the site was calculated from weather
15 variables recorded at a station located 150 m away from the experimental vineyard using the
16 Penman-Monteith equation (Allen et al., 1998). The ET₀ was then used, along with a constant
17 crop coefficient (K_c = 0.8) to compute the amount of water required by the vines (Trigo-
18 Córdoba et al., 2015). Precipitation was subtracted from ET_c each week. The calculated
19 amount of water was applied the following week.

20 Treatments consisted of a rain-fed control and an irrigation to the 50% of ET_c. Irrigation was
21 applied from late June – early July (after bloom) till mid-August, approximately two weeks
22 prior to harvest through two pressure-compensated emitters of 4 L h⁻¹ located 25 cm on either
23 side of the vine. Irrigation water was of good quality, with pH of 6.35, electrical conductivity
24 of 163.4 μS cm⁻¹ and 0.4 mg L⁻¹ of suspended solids. The water amount applied each season
25 was 40 and 50 mm for 2011 and 2012, respectively (Supplementary Table S.1).

26 **2.3 Measurements**

27 The volumetric soil water content was continuously monitored through the soil profile in two
28 spots of the experimental vineyard (one in the rain-fed treatment and another in the irrigated
29 treatment) using two capacitance probes (EnviroSCAN, Sentek, Australia), based on the

1 frequency domain reflectometry (FDR) technique. Each probe was equipped with three
2 sensors installed on an access tube at 20, 40 and 60 cm depth and connected to a datalogger.
3 The probes were properly maintained for recording soil water content at half-hour intervals
4 over the 2011 and 2012 seasons. Here, data from the irrigation period (mid-June to late-
5 August) are reported.

6 In each treatment, the probe was located within two vines (Fig. 1), avoiding to be close to the
7 emitters (25 cm from the emitter and 50 cm from the vine trunk, approximately). The
8 equation provided by the manufacturer was used for transforming permitivity data registered
9 by the probes into soil water content since we only wanted to compare relative contents
10 between these two irrigation regimes. Previous work suggest that soil type greatly affects the
11 FDR readings, but the default equation is valid for differential measurements (Paraskovas et
12 al., 2012).

13 **2.4 Multifractal analysis**

14 The concepts of multifractals and their estimation methods that were used in the current study
15 are next summarized. For detailed descriptions about multifractals, further information can be
16 found in Chhabra et al. (1989) and Everstz and Mandelbrot (1992).

17 To implement the multifractal analysis of one-dimensional soil water content time
18 distributions supported on a given interval $I = [a, b]$, a set of not-overlapping sub-intervals of I
19 with equal length is required. A common choice is to consider dyadic scaling down (Everstz
20 and Mandelbrot, 1992; Caniego et al., 2005), which means successive partitions of I in k
21 stages ($k = 1, 2, 3, \dots$). Hence, at each scale, d , a number of segments, $N(\delta) = 2^k$ are obtained
22 with characteristic time resolution, $\delta = L \times 2^{-k}$, covering the whole extent of I .

23 Multifractal approach applied to time series has already been described (Jiménez-Hornero et
24 al., 2010), hence, we only summarize the technique used in the current study. The time
25 interval of soil water content data series, L , varied from half an hour to two months and the
26 minimum time resolution, δ_{ini} , was chosen accounting for containing at least one half-hourly
27 averaged soil moisture data, θ_{ini} , at every initial interval. According to this, the probability
28 mass distribution, $p_i(\delta)$, at time resolution δ was estimated as:

$$29 \quad p_i(\delta) = \frac{\theta_i(\delta)}{\sum_j^{n_{ini}} (\theta_{ini})_j} \quad (1)$$

1 where θ_i is the water content of the i^{th} interval and n_{ini} is the number of initial intervals with
 2 mean soil water content θ_{ini} .

3 The method of the moments was used (Chhabra et al., 1989) to analyze the multifractal
 4 spectrum of the probability mass function, $p_i(\delta)$. The partition function $\chi(q, \delta)$ was estimated:

$$5 \quad \chi(q, \delta) = \sum_{i=1}^n p_i(\delta)^q \quad (2)$$

6 where moment q is a real number between $-\infty$ and $+\infty$.

7 A log-log plot of the partition function versus δ for different values of q yields:

$$8 \quad \chi(q, \delta) \propto \delta^{-\tau(q)} \quad (3)$$

9 where $\tau(q)$ is the mass scaling function of order q . The functions $f(\alpha)$ and α can be obtained by
 10 Legendre transformation of the mass exponent, $\tau(q)$, as: $f(\alpha) = \alpha(q) - \tau(q)$ and $\alpha(q) = d\tau(q)/dq$,
 11 respectively. Log-log plots of $\chi_q(\delta)$, versus δ , however, typically exhibit linearity across a
 12 limited scale range (e.g. Posadas et al., 2003), which results in drawbacks when using the
 13 moment method to obtain the singularity spectrum.

14 The direct method (Chhabra and Jensen, 1989) avoids inaccuracies associated to the
 15 estimation of $\alpha(q)$ by Legendre transformation. This method is based on the calculation of the
 16 contributions of individual segments, $\mu_i(q, \delta)$, to the partition function, which are defined as:

$$17 \quad \mu_i(q, \delta) = \mu_i^q(\delta) / \sum_{i=1}^{N(\delta)} \mu_i^q(\delta) \quad (4)$$

18 Then, using a set of real numbers, q , ($-\infty < q < \infty$), the relationships applied to calculate $f(\alpha)$
 19 and α , can be expressed as:

$$20 \quad f(\alpha(q)) \propto \frac{\sum_{i=1}^{N(\delta)} \mu_i(q, \delta) \log[\mu_i(q, \delta)]}{\log(\delta)} \quad (5a)$$

21 and

$$22 \quad \alpha(q) \propto \frac{\sum_{i=1}^{N(\delta)} \mu_i(q, \delta) \log[\mu_i(\delta)]}{\log(\delta)} \quad (5b)$$

23 The $f(\alpha)$ - α spectrum is reduced to a point for monofractal scaling type. The minimum scaling
 24 exponent (α_{min}) corresponds to the most concentrated region of the measure, and the
 25 maximum exponent (α_{max}) corresponds to the rarefied regions of the measure. A plot of $f(\alpha)$
 26 vs. α is called multifractal spectrum. It is a downward function with a maximum at $q = 0$. The

1 width of the multifractal spectrum ($w = \alpha_{max} - \alpha_{min}$) indicates overall variability (Moreno et
 2 al., 2008) similar to the nugget effects in geostatistics. For each data series, we calculated
 3 multifractal spectrum with q from -10 to $+10$ in steps of 0.5 , fine enough to show the
 4 multifractal behaviour in the studied moment range.

5 Multifractal measures can also be characterized on the basis of the generalized dimension, D_q ,
 6 of the moment of order q of a distribution, defined by Grassberger and Procaccia (1983),
 7 based on the work of Rényi (1955). The D_q , of a multifractal measure is calculated as:

$$8 \quad D_q = \frac{\tau(q)}{q-1} = \frac{1}{q-1} \lim_{\delta \rightarrow 0} \frac{\log[\chi_q(\delta)]}{\log \delta}, \quad q \neq 1 \quad (6a)$$

9 and

$$10 \quad D_1 \approx \lim_{\delta \rightarrow 0} \frac{\sum_{i=1}^{n(\delta)} \mu_i(\delta) \log[\mu_i(\delta)]}{\log \delta}, \quad q=1 \quad (6b)$$

11 Equation (6a) shows that $\tau(q)$ is also related to the generalized fractal dimension, D_q . In fact,
 12 the concept of generalized dimension, D_q , corresponds to the scaling exponent for the q^{th}
 13 moment of the measure. Using equation (6a), D_1 becomes indeterminate. Therefore, for the
 14 particular case that $q=1$, equation (6b) was employed.

15 For a monofractal, D_q is a constant function of q . However, for multifractal measures, the
 16 relationship between D_q and q is described by a S-shaped curve. In this case, the most
 17 frequently used generalized dimensions are D_0 for $q=0$, D_1 for $q=1$ and D_2 for $q=2$, which
 18 are referred to as capacity, information (or Shannon entropy) and correlation dimension,
 19 respectively. The information dimension, D_1 , provides insight about the degree of
 20 heterogeneity in the distribution of the measure. The correlation dimension, D_2 , is associated
 21 to the uniformity of the measure among intervals and describes the average distribution
 22 density of the measure. In general, the generalized dimension, D_q , is more useful for the
 23 comprehensive study of multifractals. Differences between D_q allow comparison of the
 24 complexity between measured soil water content data series. In homogeneous structures D_q
 25 are close, whereas in a monofractal they are equal.

26

1 **3 Results and discussion**

2 **3.1 Patterns of vineyard soil water content under rain-fed and irrigation** 3 **conditions**

4 Temperatures for the two studied growing seasons were similar in average (Table 1);
5 however, rainfall and evapotranspiration were higher in 2012. Harvest date was almost the
6 same in both years. Nevertheless, the temporal evolution of rainfall and ET_c differed from
7 year to year (Fig. 2), being greater during 2012, especially at the beginning of the study
8 period. This fact caused a different scheduling of irrigation between years.

9 Soil water content decreased over the growing season under rain-fed conditions in both years
10 (Fig. 3). However, when irrigation was initiated, soil water content became more stable in the
11 irrigated treatment (Fig. 3). The magnitude of the soil water loss was more evident in the
12 layers of 20 and 40 cm depth, and less important in the 60 cm layer, which may indicate the
13 depth of the active root zone as well as the intensity of root water uptake at each soil layer, as
14 reported for other cultivars and crops (Intrigliolo and Castel, 2009; Mestas-Valero et al.,
15 2011), and proved that FDR probes can be successfully used for irrigation scheduling
16 (Goldhamer et al., 1999), calibrating them with established indicators such as midday stem
17 water potential (Mirás-Avalos et al., 2014) and soil evaporation..

18 **3.2 Multifractality of the soil water content time series**

19 Soil water content time series obeyed power law scaling, as shown by the double log plots
20 (Supplementary Fig. S.1). These plots allow to identify the range of moments needed to
21 describe the scale variation of the studied parameter (Vidal Vázquez et al., 2010).

22 Figure 4 shows the partition functions for rain-fed and irrigation conditions at 20 cm depth in
23 2011. Visually, a slight departure from the straight line model was observed for moments $q <$
24 -1 (Supplementary Fig. S.1). In general, higher deviations from linearity were found for the
25 highest q moments in the data series from the irrigation treatment, when compared to those
26 from the rain-fed treatment, especially in 2012. Nevertheless, determination coefficients, R^2 ,
27 were greater than 0.9 for statistical moments in the range from $q = -10$ to $q = 10$, in all the
28 studied data sets. Consequently, scalings are adequately defined. Similar results were found
29 by Mestas-Valero et al. (2011) for soil water content under rain-fed grassland.

1 The $\tau(q)$ functions were different from a monofractal type of scaling for all series analyzed,
2 especially under irrigation conditions (Supplementary Fig. S.2), similar to results obtained by
3 Biswas et al. (2012b) for soil water storage. In fact, the heterogeneity of the soil water
4 content data series from the irrigated treatment was greater than that of the rain-fed treatment
5 (Supplementary Fig. S.2).

6 The value of D_I is a good indicator of the heterogeneity degree in temporal distributions of a
7 given variable. The closer the D_I value to D_0 , the more homogeneous is the distribution of the
8 variable. In our case, rain-fed series were more homogeneous than the irrigated ones. In
9 general, soil water content recorded at 60 cm depth presented the lower differences between
10 D_I and D_0 (Table 1), thus being more homogeneous both under rain-fed and irrigation
11 conditions. Moreover, 2012 data series presented a higher heterogeneity than those from
12 2011 (Table 1) for both treatments, caused by the greater rainfall amount collected in 2012.

13 A monofractal would be characterized by $D_0 = D_I = D_2$ (Evertsz and Mandelbrot, 1992). In all
14 the studied data series $D_0 > D_I > D_2$ (Table 1), indicating that soil water content had a
15 tendency to behave as a multifractal. However, differences ($D_0 - D_I$) ranged from 0.051 to
16 0.222 and ($D_I - D_2$) oscillated between 0.053 and 0.168, which suggests different degree in
17 the homogeneity/heterogeneity of soil water content depending on the treatment imposed and
18 the depth in the soil profile. In general, data series from the irrigation treatment showed
19 greater differences between D_0 , D_I and D_2 than the series from the rain-fed treatment for both
20 growing seasons. Moreover, the 60 cm depth layer presented smaller differences than the 20
21 and 40 cm layers (Table 1). The width of the D_q spectra, determined by indicators such as (D_0
22 $- D_{10}$), showed different degrees of heterogeneity, with a trend to decrease in depth and
23 under rain-fed conditions when compared with the irrigation treatment (Table 1). This is
24 caused by the spiky nature of soil water content and indicates a multiple scaling nature at
25 shallow depths. Moreover, the width of the D_q spectra increased from 2011 to 2012 in both
26 treatments, mainly in the 20 and 40 cm depths.

27 Generalized dimensions, or Rényi spectra, calculated for the range between $q = -10$ and $q =$
28 10 for soil water content data series at three depths under rain-fed and irrigation conditions
29 are displayed on Fig. 4. All the data series studied showed Rényi spectra as asymmetric
30 sigma-shaped curves with more curvature for the negative values of q than for positive ones
31 (Fig. 4). The left part of the curves is concave down and it changes to concave up on the right
32 of the vertical axis. In the case of the soil water content series from the rain-fed treatment, the

1 most curved spectra corresponded to the 40 cm depth data series, whereas for the irrigation
2 treatment, the most curved one was the 20 cm depth data series (Fig. 4). When compared
3 between treatments, Rényi spectra were more curved under irrigation conditions and the
4 estimation errors were also greater under this treatment (Fig. 4). These results confirmed the
5 higher heterogeneity (multifractality) of the data series from the irrigation treatment when
6 compared to those from rain-fed.

7 Mestas-Valero et al. (2011) obtained monofractal distributions of soil water content time
8 series under grassland when measured at depths greater than 40 cm, in contrast with our
9 results. This disagreement is likely caused by the fact that grapevine root system reach greater
10 depths than that of grass and vines are capable of uptaking water from deeper soil layers.

11 Determination coefficients, R^2 , were highest for moments $q = 0$ and $q = 1$ and diminished for
12 the other $|q|$ moments. In the case of $q = 10$, R^2 was greater than 0.97 and 0.95 in the rain-fed
13 and irrigated data sets, respectively. For $q = -10$, R^2 values for rain-fed and irrigated data
14 series were greater than 0.99 and 0.91, respectively (data not shown). Standard errors of D_q
15 values increased with increasing $|q|$ moments and they were much lower for right ($q > 0$) than
16 for left ($q < 0$) branch of the Rényi spectra (Fig. 4).

17 Parameter α_0 from the singularity spectra ranged from 1.056 to 1.146 in the rain-fed treatment
18 and from 1.075 to 1.187 in the irrigated treatment (Table 2). The singularity spectrum allows
19 for analyzing similarity or difference between the scaling properties of the measures as well
20 as to assess the local scaling properties of soil water content measurements. The wider the
21 spectrum is (i.e., the largest $\alpha_{q-} - \alpha_{q+}$ value), the higher the heterogeneity in the scaling
22 indices and vice versa (Vidal Vázquez et al., 2010). Moreover, the $f(\alpha)$ spectrum branch
23 length gives insight about the abundance of the measure. Hence, small $f(\alpha)$ values at the end
24 of a long branch correspond to rare events. Our results showed that the width of the
25 singularity spectra increased in both treatments from 2011 to 2012 (Table 2).

26 Singularity spectra are characterized by a concave down shape (Fig. 5), showing an
27 asymmetrical curve with wider but shorter right side. Rain-fed data series showed a shorter
28 $f(\alpha)$ spectrum in both years, confirming their low degree of multifractality when compared to
29 the irrigated data series (Fig. 5).

30 Differences ($\alpha_{q-} - \alpha_0$ and $\alpha_0 - \alpha_{q+}$) indicate the deviation of the spectrum from its maximum
31 value ($q = 0$) towards the right side ($q < 0$) and the left side ($q > 0$), respectively (Vidal

1 Vázquez et al., 2010). Usually, soil water content data series from the rain-fed treatment
2 showed lower $\alpha_0 - \alpha_{q+}$ values than those from the irrigated treatment (Table 2). Moreover, the
3 highest values for this multifractal parameter were observed at 40 cm depth in both treatments
4 and years (Table 2). This may indicate that higher soil water contents were more frequent
5 under irrigation, being greater the differences between treatments at 40 cm depth in 2012. In
6 contrast, the right branch ($\alpha_{q-} - \alpha_0$) of the spectrum was usually wider for rain-fed conditions
7 (Table 2). These results confirm the differential homogeneity/heterogeneity pattern between
8 treatments evidenced by the generalized dimension, D_q , analysis (Table 1, Fig. 4).

9

10 **4 Conclusions**

11 Under the conditions of this study, continuous soil water content measurements at different
12 depths reliably described the soil water balance in a vineyard over two irrigation periods.

13 The logarithms of the partition function varied linearly with the logarithms of the time
14 resolution for all the studied depths under both treatments considered in the range of moments
15 $-10 < q < 10$, indicating that soil water content time series obeyed power laws.

16 The scaling properties of soil water content time series were reasonably fitted to multifractal
17 models. These properties were different for the rain-fed and irrigation treatments, implying a
18 higher heterogeneity for the data series from the irrigation treatment, which tended to increase
19 in the second year of the study (2012). Therefore, multifractal analysis allowed us to
20 discriminate among soil water content patterns in a vineyard for the 2011 and 2012 growing
21 seasons as a function of irrigation use.

22

23 **Author contribution**

24 J. M. Mirás-Avalos and E. Trigo-Córdoba designed and carried out the field experiment. J. M.
25 Mirás-Avalos, R. da Silva-Dias, I. Varela-Vila and A. García-Tomillo performed the
26 analyses. J. M. Mirás-Avalos prepared the manuscript with contributions from all co-authors.

27

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3

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1 Table 1. Selected multifractal parameters: generalized dimensions, for the first-three positive
 2 moments, D_0 , D_1 , and D_2 , with their respective errors of estimation, and two multifractality
 3 indices $\Delta(D_0 - D_2)$ and $\Delta(D_0 - D_{10})$.

Treatment	Depth (cm)	D_0	D_1	D_2	$\Delta(D_0 - D_2)$	$\Delta(D_0 - D_{10})$
2011						
Rain-fed	20	0.999 ± 0.001	0.937 ± 0.008	0.884 ± 0.016	0.115	0.672
	40	1.000 ± 0.000	0.881 ± 0.007	0.746 ± 0.014	0.254	0.752
	60	1.000 ± 0.000	0.925 ± 0.007	0.868 ± 0.013	0.133	0.656
	20-60	1.000 ± 0.000	0.916 ± 0.008	0.833 ± 0.019	0.167	0.589
Irrigated	20	0.999 ± 0.001	0.868 ± 0.013	0.778 ± 0.026	0.221	0.757
	40	1.000 ± 0.000	0.852 ± 0.019	0.773 ± 0.026	0.227	0.698
	60	1.000 ± 0.000	0.852 ± 0.022	0.758 ± 0.034	0.242	0.664
	20-60	1.000 ± 0.000	0.861 ± 0.023	0.773 ± 0.037	0.227	0.695
2012						
Rain-fed	20	0.999 ± 0.001	0.861 ± 0.014	0.771 ± 0.025	0.228	0.856
	40	1.000 ± 0.000	0.888 ± 0.008	0.739 ± 0.017	0.261	0.801
	60	1.000 ± 0.000	0.949 ± 0.004	0.907 ± 0.005	0.093	0.548
	20-60	1.000 ± 0.000	0.898 ± 0.006	0.768 ± 0.016	0.232	0.682
Irrigated	20	0.984 ± 0.006	0.831 ± 0.010	0.731 ± 0.019	0.253	1.024
	40	0.979 ± 0.006	0.757 ± 0.014	0.589 ± 0.022	0.390	1.210
	60	1.000 ± 0.000	0.907 ± 0.007	0.805 ± 0.015	0.195	0.622
	20-60	0.993 ± 0.003	0.822 ± 0.016	0.707 ± 0.030	0.286	1.085

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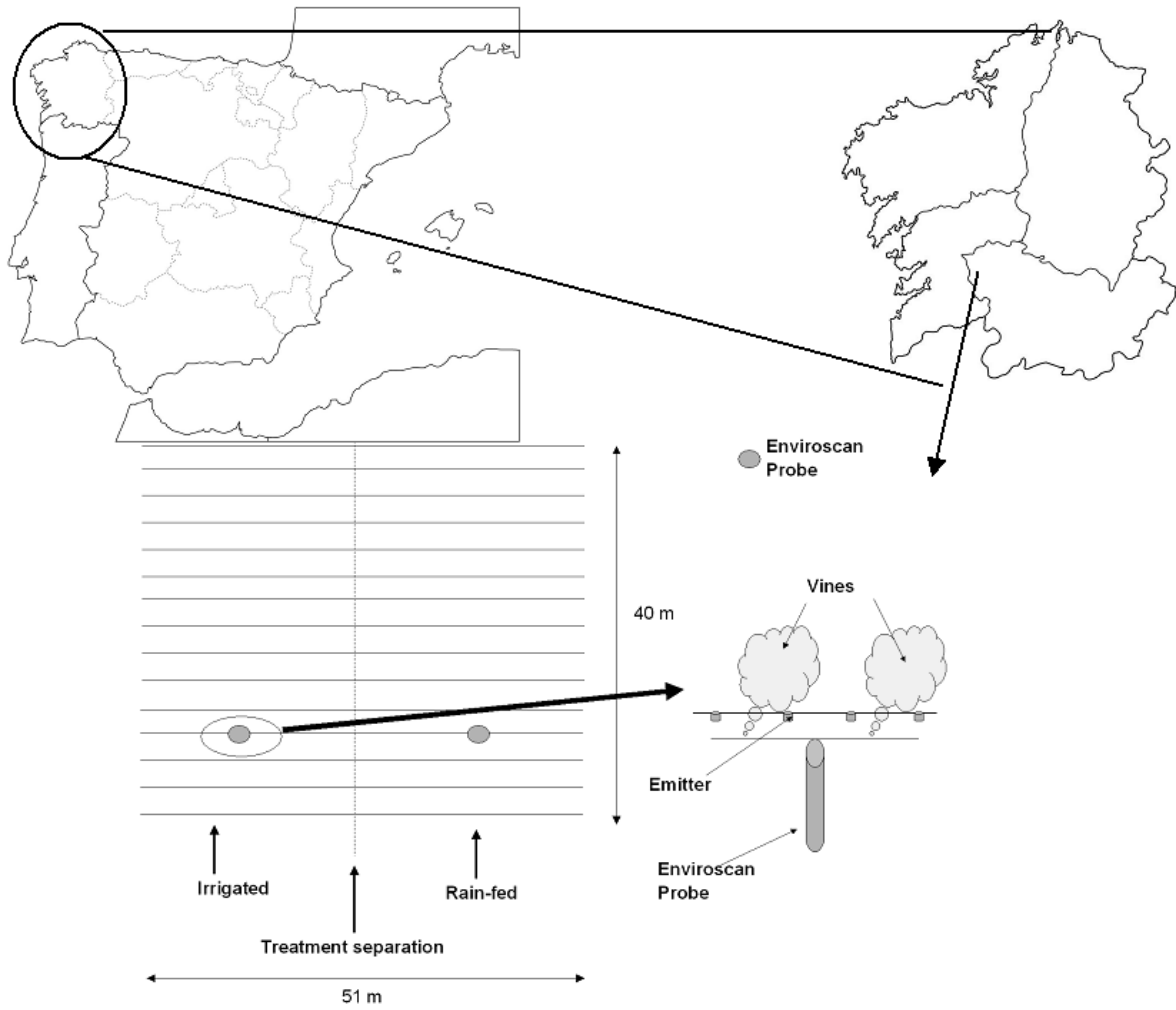
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1 Table 2. Selected multifractal parameters derived from the $f(\alpha)$ singularity spectra: most
 2 positive (q_+) and most negative (q_-) limits the range of multifractal scaling, Hölder exponent
 3 of order 0 (α_0), most positive (α_{q_+}) and most negative (α_{q_-}) exponents, widths of the left ($\alpha_0 -$
 4 α_{q_+}) and the right ($\alpha_{q_-} - \alpha_0$) sides of the spectra.

Treatment	Depth (cm)	q_-	q_+	α_0	α_{q_+}	α_{q_-}	$\alpha_0 - \alpha_{q_+}$	$\alpha_{q_-} - \alpha_0$
2011								
Rain-fed	20	-1.5	3.5	1.066	0.768	1.339	0.299	0.273
	40	-3.5	2	1.093	0.632	1.328	0.460	0.235
	60	-3.5	2	1.087	0.718	1.403	0.369	0.315
	20-60	-4	2	1.074	0.762	1.297	0.312	0.222
Irrigated	20	-2.5	2	1.136	0.714	1.450	0.422	0.314
	40	-4	3	1.160	0.664	1.383	0.496	0.222
	60	-5	2	1.132	0.700	1.333	0.435	0.200
	20-60	-4.5	2	1.142	0.709	1.375	0.433	0.233
2012								
Rain-fed	20	-2.5	3	1.146	0.659	1.526	0.487	0.380
	40	-3.5	2	1.082	0.603	1.301	0.479	0.219
	60	-2	5.5	1.056	0.746	1.296	0.309	0.240
	20-60	-5	2	1.077	0.651	1.265	0.426	0.188
Irrigated	20	-0.5	2.5	1.164	0.602	1.361	0.562	0.197
	40	-1	1.5	1.187	0.575	1.491	0.611	0.304
	60	-4	2	1.075	0.716	1.223	0.360	0.148
	20-60	-1	2	1.172	0.624	1.489	0.548	0.317

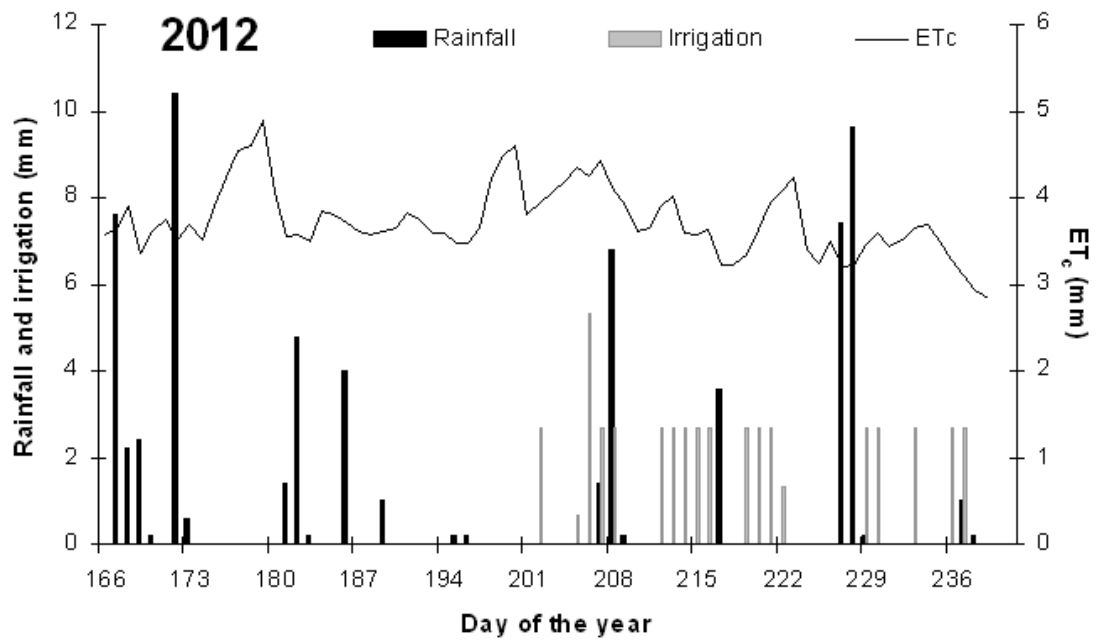
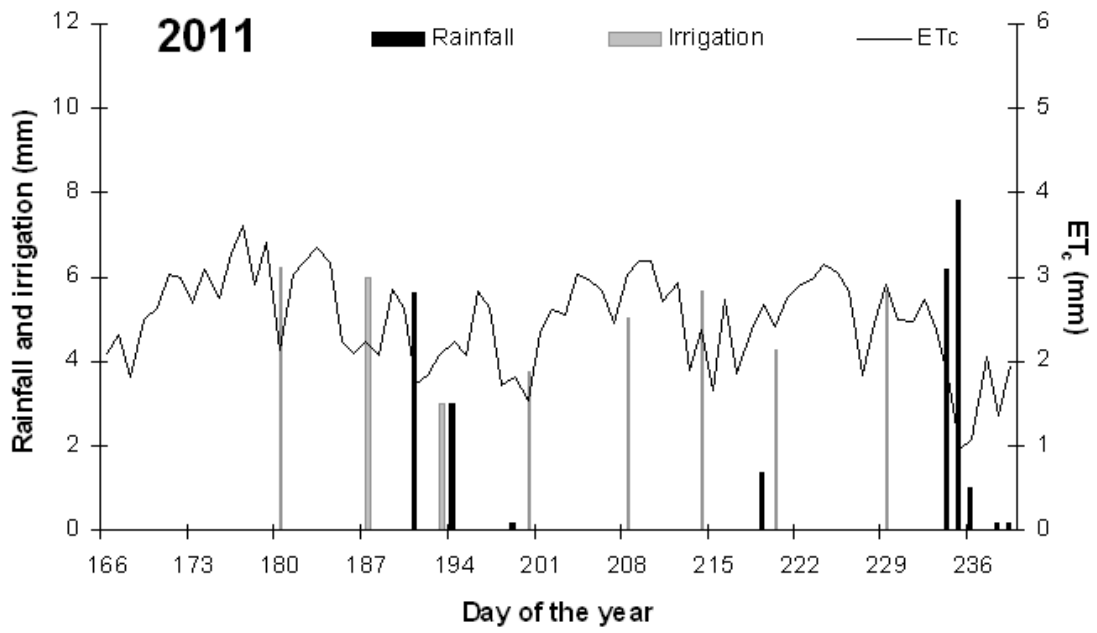
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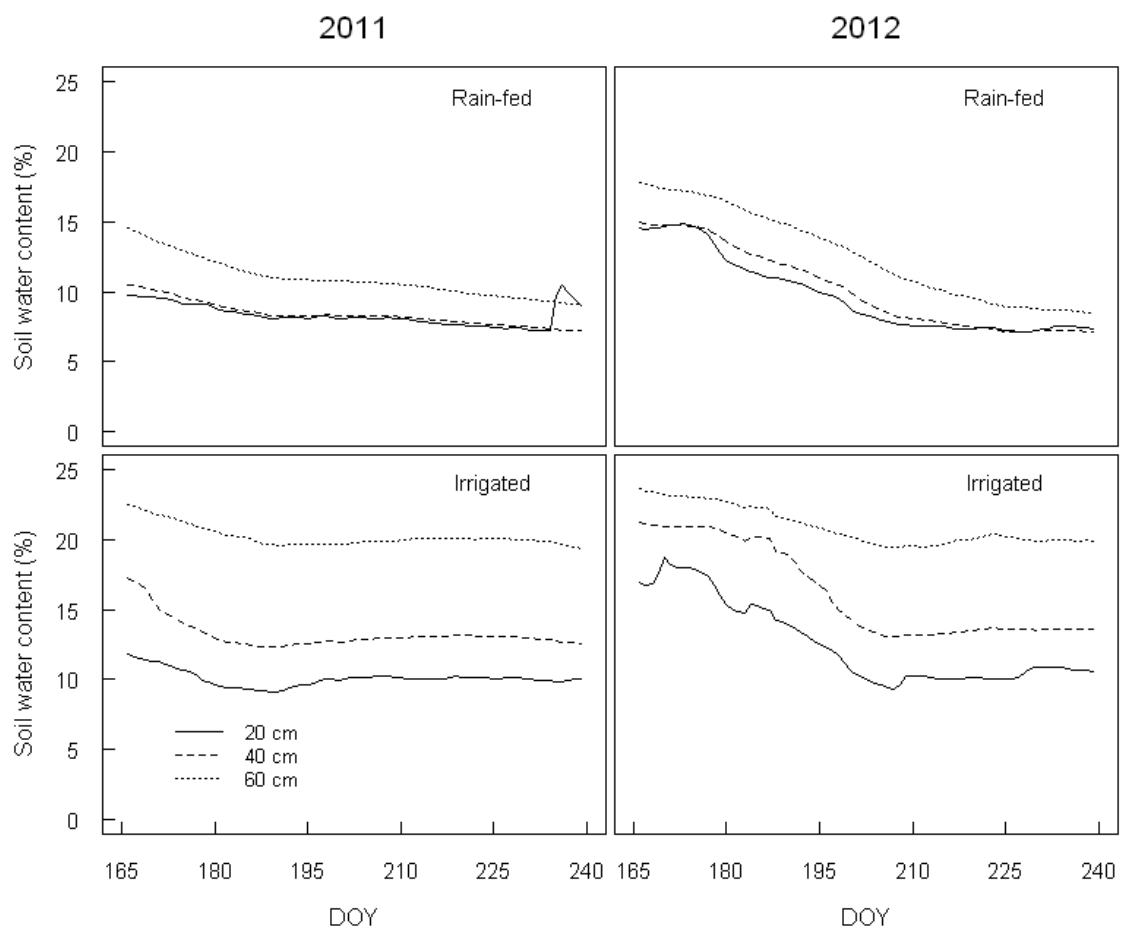
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Figure 1. Location of the studied vineyard and experimental layout.



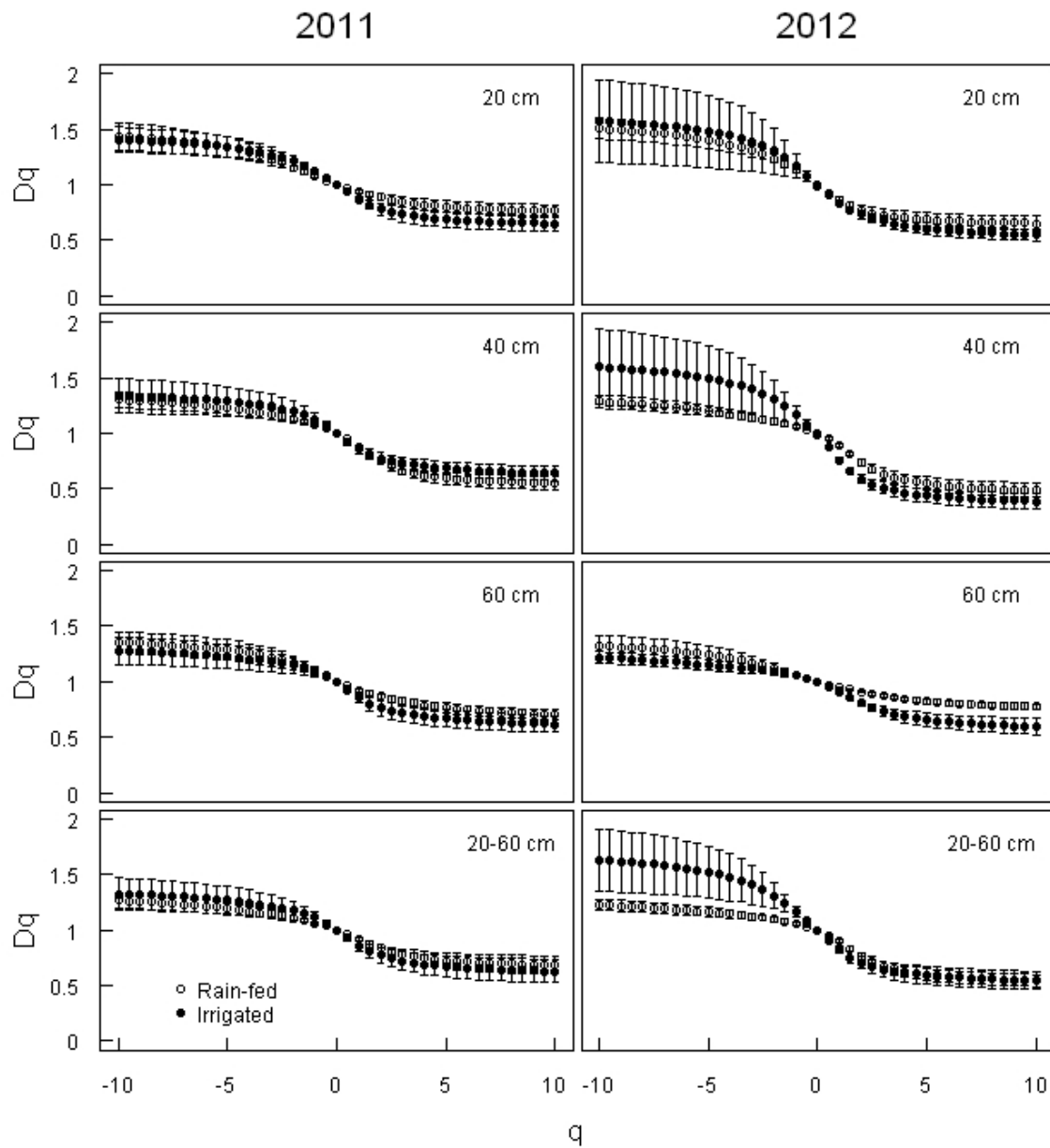
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Figure 2. Crop evapotranspiration (ET_c), rainfall and irrigation water applied over the two growing seasons studied, 2011 and 2012. Day of the year 166 is 14th June.



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 2 Figure 3. Soil water content at three depths (20, 40 and 60 cm) for rain-fed and irrigation
 3 treatments over the 2011 and 2012 growing seasons. DOY stands for Day of the Year (165 =
 4 13th June).

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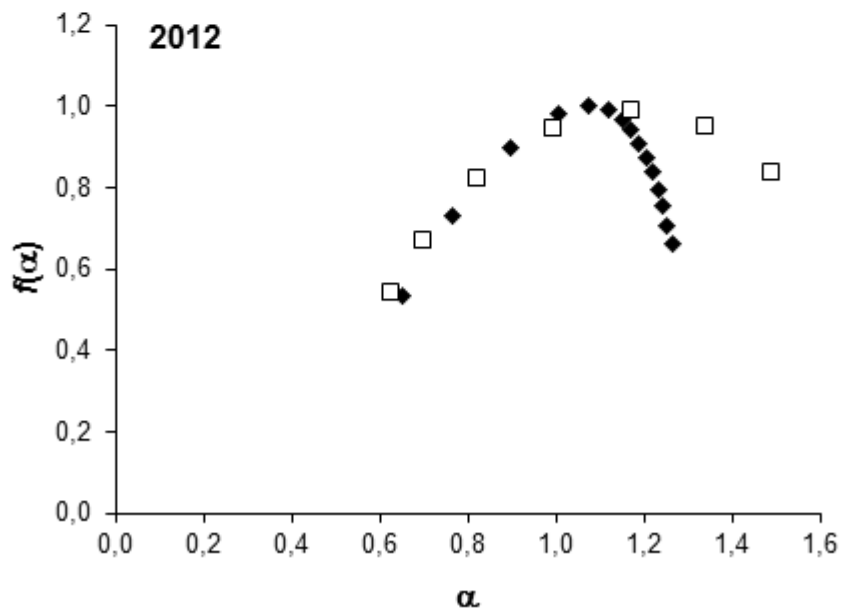
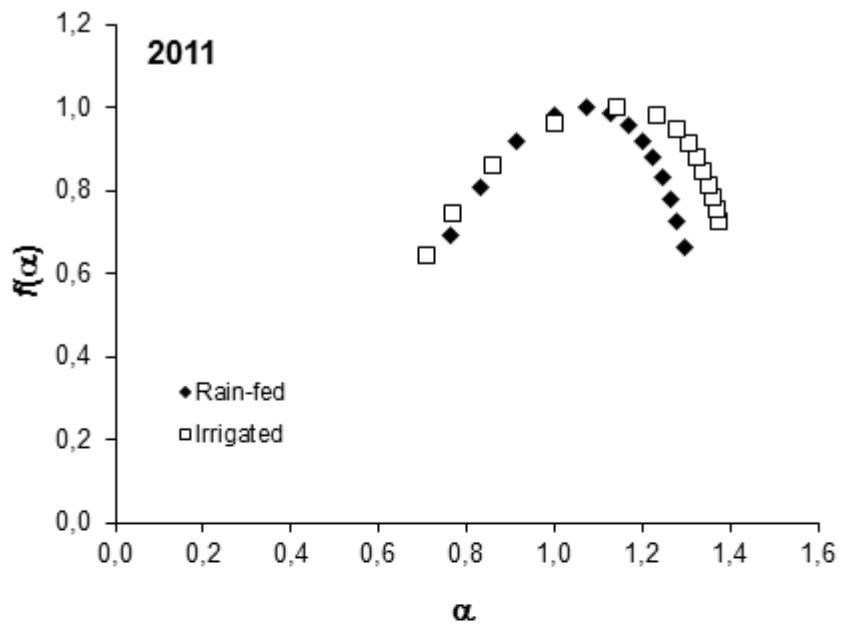
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3 Figure 4. Generalized dimension, D_q , spectra ($-10 < q < 10$) of soil water content for rain-fed
 4 and irrigation treatments at the studied depths in 2011 and 2012. Bars indicate estimation
 5 errors.

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3 Figure 5. Singularity spectra for soil water content averaged from 20 to 60 cm depth for rain-
4 fed and irrigation treatments in 2011 and 2012.